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# Weights in CPI/HICP and in seasonally adjusted series

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#### Remarks:

Part of the research for this paper was done when the author was working as a Seconded National Expert in Eurostat. The views expressed in this paper are those of the author and do not necessarily reflect the policies of Eurostat or Statistics Netherlands.

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## WEIGHTS IN CPI/HICP AND IN SEASONALLY ADJUSTED SERIES

*Summary: In this paper we discuss the weights in a Consumer Price Index (CPI)<sup>1</sup>. Both the fixed base Laspeyres index and chain-linked Laspeyres-type indices like the European Harmonised Index of Consumer Prices (HICP) are discussed. How are the weights derived and what data are usually published. The central question is what should be the weights in the case where seasonally adjusted underlying series are aggregated to get a seasonally adjusted headline index series. In the case of a chain-linked index the optimum weights differ from the published ones. We discuss the possibilities to derive the optimum weights and give a numerical example.*

*Keywords:* CPI, HICP, Laspeyres type index, weights, Seasonal adjustment

### 1. Introduction

Consumer prices often have a distinct seasonal pattern. The most widely used response to the seasonal variation of the Consumer Price Index (CPI) is to focus on year-on-year inflation results. Another response is to calculate seasonally adjusted indices and month-on-month rates of change that reflect recent developments in inflation which abstract from regular seasonal effects.

Many statistical institutes, including Statistics Netherlands and Eurostat, do not produce seasonally adjusted series for the CPI or HICP (European Harmonised Index of Consumer Prices) themselves but other organisations may use official statistics to produce them.

Basically two approaches for compiling these series are possible. In a “direct approach” seasonal adjustment is applied directly to each series of interest, e.g. the headline overall index series. There is also an “indirect approach”. Considering that the overall CPI is the weighted aggregate of a number of underlying price index series, it is also possible to apply seasonal adjustment to (a subset of) these underlying series and then aggregate them to get a seasonally adjusted overall aggregate index series.

This paper deals with indirect seasonal adjustment methods for the CPI. It is attempted to answer the question which weights must be used for aggregating the seasonally adjusted elementary aggregates in order to get an unbiased seasonally

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<sup>1</sup> The author wishes to thank Martin Eiglsperger (ECB) for valuable comments on an earlier draft. All remaining errors are mine.

adjusted headline series. It is not about the seasonal adjustment of the elementary aggregates itself.

In section 2 we discuss the alternative ways to represent a fixed base Laspeyres index and the corresponding weights. This leads in section 3 to a more general representation of a chain-linked Laspeyres index, which is used for the HICP. In section 4 we discuss the combining of weights based on annual consumption with monthly indices and the related problems of seasonalities. Section 5 presents the weights in an annually chained CPI with a monthly index. In section 6 we give a short description of the HICP-regulation on weights and the weights that are actually published by Eurostat.

In section 7 the direct and indirect approach for seasonal adjustment are compared. The rest of the paper will be on the indirect approach. In section 8 we discuss the weights to be used in a seasonally adjusted fixed base Laspeyres series and in section 9 we do the same for a seasonally adjusted chain-linked series like the HICP. In the latter case the published weights are not the best for the seasonally adjusted series on theoretical considerations. It is also derived how corrected weights for seasonally adjusted series can be derived from the published weights.

In section 10 we make a comparison of results on the basis of seasonally adjusted HICP series published by the European Central Bank (ECB). In the presented examples the differences between the alternative results are small, but we will discuss under what circumstances these differences may be larger.

## 2. Weights in a Laspeyres type CPI

### 2.1 The basic Laspeyres formula

This paper deals with CPIs that are compiled as Laspeyres type index series. A Laspeyres price index can be described fully by prices and quantities of a basket of goods and services, purchased by consumers.

Let  $P_{ia}$  and  $Q_{ia}$  be the price and quantity purchased of good or service  $i$  in period  $a$ .

The Laspeyres index  $U_{ab}$  measures the average price development between periods  $a$  and  $b$  of a basket of all goods and services  $i$  that were purchased in period  $a$ . It can be written as:

$$(1) \quad U_{ab} = \frac{\sum_i P_{ib} * Q_{ia}}{\sum_i P_{ia} * Q_{ia}}$$

The formula for the overall price index is usually transformed into a formula that separates the overall index into the price change of each product (a good or a service) and related weight indicators.

$$(2) \quad U_{ab} = \frac{\sum_i P_{ib} * Q_{ia}}{\sum_i P_{ia} * Q_{ia}} = \frac{\sum_i P_{ib} / P_{ia} * P_{ia} * Q_{ia}}{\sum_i P_{ia} * Q_{ia}} = \sum_i \frac{P_{ia} * Q_{ia}}{\sum_i P_{ia} * Q_{ia}} * \frac{P_{ib}}{P_{ia}} = \sum_i W_{ia} * \frac{P_{ib}}{P_{ia}}$$

with

$$(3) \quad W_{ia} = \frac{P_{ia} * Q_{ia}}{\sum_i P_{ia} * Q_{ia}}$$

The weights  $W_{ia}$  represent the expenditure shares of products  $i$  in base period  $a$ . This is also the formula normally used in practical calculations. The weights  $W_{ia}$  and elementary price indices  $P_{ib}/P_{ia}$  are not observed for all individual products but at some elementary product group level.

## 2.2 A chain linked representation of the fixed basket Laspeyres index.

A Laspeyres index covering a longer period can also be written as a chain linked index, even if the basket of goods and services does not change. Each link is a Lowe index<sup>2</sup> covering shorter periods.

$$(4) \quad U_{ab} = \frac{\sum_i P_{ib} * Q_{ia}}{\sum_i P_{ia} * Q_{ia}} = \frac{\sum_i P_{i(a+1)} * Q_{ia}}{\sum_i P_{ia} * Q_{ia}} * \frac{\sum_i P_{i(a+2)} * Q_{ia}}{\sum_i P_{i(a+1)} * Q_{ia}} * \dots * \frac{\sum_i P_{ib} * Q_{ia}}{\sum_i P_{i(b-1)} * Q_{ia}}$$

This can be written as

$$(5) \quad U_{ab} = \prod_{x=a}^{b-1} \left( \frac{\sum_i P_{i,x+1} * Q_{ia}}{\sum_i P_{ix} * Q_{ia}} \right) = \prod_{x=a}^{b-1} \left( \sum_i W_{i,a,x} * \frac{P_{i,x+1}}{P_{ix}} \right)$$

with

$$(6) \quad W_{i,a,x} = \frac{P_{ix} * Q_{ia}}{\sum_i P_{ix} * Q_{ia}} = W_{i,a,a} * \frac{P_{ix}}{P_{ia}} * \frac{\sum_i P_{ia} * Q_{ia}}{\sum_i P_{ix} * Q_{ia}}$$

In this case  $a$  represents the base period in which the basket of goods and services was determined and  $x$  the index reference period (period where the index=1) for the chain-link.

Note that in this chain-link representation of the fixed basket Laspeyres index the sets of weights  $W_{i,a,x}$  are calculated in each period using the same product quantities  $Q_{ia}$  but they differ because of different price relatives of the products in each period. This right hand term of equation 6 shows the process of price-updating. The three parts of the expression on the right of equation 6 represent:

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<sup>2</sup> The Lowe index is a more general index than the Laspeyres index in that the quantities  $Q_{ia}$  in a Laspeyres index represent the consumption basket in the base period  $a$ , whereas in the Lowe index the period in which the consumption basket was measured is not strictly defined.

- the weights in the base year at base year prices,
- the relative of the price of product  $i$  in period  $x$  compared to base period  $a$ ,
- the inverse of the price development of the full basket between the two periods.

### 2.3 The general case of the chain-linked Laspeyres index

A more general case of the Laspeyres formula for a longer term period is one where the weights may be adjusted on other grounds than changing price relatives, namely changing consumption patterns. The weights  $W_{ix}$  in equation (7) represent estimates for the expenditure shares of product  $i$  used in the index calculation in period  $x+1$ .

$$(7) U_{ab} = \prod_{x=a}^{b-1} \left( \sum_i W_{ix} * \frac{P_{i,x+1}}{P_{ix}} \right)$$

Note that the subscript  $a$  has disappeared in the weights since the weights in period  $x$  are no longer based on the expenditures in period  $a$ . Each period a new estimate for the weights is entered and used for the calculation of the index in  $t+1$ .

## 3. Annual weights and monthly indices

The consumption basket in the CPI is usually based on a 12-month consumption period. This is very often, but not necessarily a calendar year. The choice of a 12-month period is based on the fact that the consumption pattern is not the same in each month of the year, due to varying weather conditions, availability of seasonal products, holiday periods and etcetera. By taking the consumption in 12 months (or a multiple of 12 months) we make sure that all consumption during the year is covered by the CPI.

This does not prevent statisticians from making monthly or quarterly CPIs. The production of monthly figures may in principle introduce two kinds of seasonality: seasonality in the consumption pattern and seasonality in the prices. The aim of the CPI is to measure the development of prices and therefore the impacts of changing consumption patterns are eliminated to the extent possible.

Seasonality in consumption patterns is related to seasonal products that are not available for purchase by consumers in all months. This also means that prices for these products cannot be observed when they are out of season. The HICP-regulations<sup>3</sup> have defined two distinct methods that can be used for the treatment of seasonal products.

<sup>3</sup> Details of the HICP methods can be found in the Commission Regulation (EC) no. 330/2009 of 22 April 2009. It goes beyond the scope of this paper to discuss fully the treatment of seasonal products.

The first one is the “class confined seasonal weights method”. This method is based on seasonal product weights, with restrictions that ensure that product weights add up to fixed aggregate weights at each subdivision of COICOP.

The second allowed method is the use of a strict annual weights index where missing prices are estimated using either “counter-seasonal estimation” or “all-seasonal estimation”.

What is important here is that in both methods for the treatment of seasonal products in the HICP, the COICOP weights are fixed at each subdivision of COICOP.

The fact that in principle there is no seasonal pattern in the consumption as represented by the weighting scheme of the CPI reflects the general property of the CPI that the monthly indices do not represent the prices for the basket consumed in the reporting month but the price level of the annual basket in that month.

For the rest of this paper we will only deal with the seasonal patterns of prices, not with the seasonal patterns of consumption.

#### 4. Weights in a chain-linked and monthly HICP

The HICP is a chain-linked Laspeyres-type price-index series<sup>4</sup>. For a month  $m$  in the period of 12 months between December  $t-1$  and December  $t$  the index is calculated and linked as in equations (8) and (9). The index reference period or the base year of the published series is 2005. In these equations  $m$  represents the reporting month and  $D$  represents December:

$$(8) U_{2005(m,A)} = \frac{\sum_i P_{i(D,2005)} * Q_{i,2005}}{\sum_i P_{i,2005} * Q_{i,2005}} * \prod_{x=2005}^{A-2} \left( \frac{\sum_i P_{i(D,x+1)} * Q_{i(D,x)}}{\sum_i P_{i(D,x)} * Q_{i(D,x)}} \right) * \frac{\sum_i P_{i(m,A)} * Q_{i(D,A-1)}}{\sum_i P_{i(D,A-1)} * Q_{i(D,A-1)}}$$

This can be written as:

$$(9) U_{2005(m,A)} = \left( \sum_i W_{i(2005)} * \frac{P_{i(D,2005)}}{P_{i(2005)}} \right) * \prod_{x=2005}^{A-2} \left( \sum_i W_{i(D,x)} * \frac{P_{i(D,x+1)}}{P_{i(D,x)}} \right) * \left( \sum_i W_{i(D,A-1)} * \frac{P_{i(m,A)}}{P_{i(D,A-1)}} \right)$$

with

$$(10) W_{i(D,x)} = \frac{P_{i(D,x)} * Q_{i(D,x)}}{\sum_i P_{i(D,x)} * Q_{i(D,x)}}$$

<sup>4</sup> The term "Laspeyres-type" follows from the HICP-regulations of 1996. Actually the HICP is a chain-linked series of Lowe indices. The use of the term Lowe index was introduced with the ILO-CPI-manual in 2004.

is the December weight for product group  $i$  in the year  $x$ . This weight is used from December  $x$  till December  $x+1$ . Furthermore

$$(11) \quad W_{i(2005)} = \frac{P_{i(2005)} * Q_{i(2005)}}{\sum_i P_{i(2005)} * Q_{i(2005)}}$$

is the annual weight for 2005, the index reference year for the long time HICP series.<sup>5</sup>

It was explained in section 4 that the monthly indices are based on annual baskets of goods and services and therefore  $Q_{i(D,A)}$  is in fact  $Q_{iA}$  and equation (10) may be rewritten as

$$(12) \quad W_{i(D,A)} = \frac{P_{i(D,A)} * Q_{iA}}{\sum_i P_{i(D,A)} * Q_{iA}} = W_{i,A} * \frac{P_{i(D,A)}}{P_{iA}} * \frac{\sum_i P_{iA} * Q_{iA}}{\sum_i P_{i(D,A)} * Q_{iA}}$$

where

$$(13) \quad W_{i,A} = \frac{P_{iA} * Q_{iA}}{\sum_i P_{iA} * Q_{iA}}$$

Equation (12) states that the weight in December of year  $A$  for product group  $i$  can be written as the product of three elements<sup>6</sup>:

- the weight in the full year  $A$ .

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<sup>5</sup> Actually the HICP series started in 1996. In 2006 the results at all levels of publication were rescaled to make the average indices in 2005 equal 100.

<sup>6</sup> Note that  $P_{iA}$  is not the average transaction price in year  $A$  but the average of the 12 monthly price indices for product group  $i$ :

$$(5.1) \quad P_{iA} = 1/12 * \sum_{m=1}^{12} P_{i(m,A)}$$

The average expenditure share for a product  $i$  in year  $A$  could be written as

$$(5.2) \quad W_{i,A} = \frac{\sum_{m=1}^{12} P_{i(m,A)} * Q_{i(m,A)}}{\sum_i \sum_{m=1}^{12} P_{i(m,A)} * Q_{i(m,A)}}$$

However, since in the CPI  $Q_{i(m,A)}=Q_{iA}$  we can write

$$(5.3) \quad W_{i,A} = \frac{\sum_{m=1}^{12} P_{i(m,A)} * Q_{iA}}{\sum_i \sum_{m=1}^{12} P_{i(m,A)} * Q_{iA}} = \frac{Q_{iA} * \sum_{m=1}^{12} P_{i(m,A)}}{\sum_i Q_{iA} * \sum_{m=1}^{12} P_{i(m,A)}} = \frac{Q_{iA} * P_{iA}}{\sum_i Q_{iA} * P_{iA}}$$

- The price relative for product group  $i$  comparing the December price index with the annual average price index
- The inverse of the Laspeyres price index comparing December prices of the full basket of year  $A$  with the annual average prices in year  $A$ .

## **5. The relationship between the consumption basket and the published weights**

Consumption patterns change over time and therefore a regular update of the weighting schemes of the CPI is necessary to guarantee the representativeness of the basket. A decade ago base revisions once every five years were rather common, but ever more countries have changed to an annual base revision. In the same period there was a development of the sources used for the weights. Where in the past the Household Budget Survey was the main source for the weights, the focus has shifted towards weights based on National Accounts consumption data.

The HICP regulation on weights prescribes that weights in the HICP in year  $t$  are based on an estimate of the consumption pattern of year  $t-1$ . When the weights for year  $t$  are first needed, in January of year  $t$ , National Accounts (NA) results for the year  $t-1$  are not yet available. Therefore NA-consumption data for the year  $t-2$  are used. Under normal circumstances the distribution of expenditures in  $t-2$  is used directly as an estimate for the expenditures in year  $t-1$ .<sup>7</sup>

Member states are recommended to check the weights and correct for known sudden changes in the expenditures distribution for year  $t-1$  before finalizing the year  $t-1$  HICP annual weights. In a final step these annual weights are price-updated to December  $t-1$  weights, i.e the expenditure shares estimated for year  $t-1$  are expressed in prices of December of that year  $t-1$ .

## **6. Seasonal adjustment; direct and indirect approach**

Seasonality in prices complicates the interpretation of short term index development. These problems can be solved either by focussing on year-on-year changes of the index, the annual inflation, or by seasonal adjustment of the series.

The headline index series is an aggregate of the price indices of underlying series. Seasonal adjustment of the headline series can be performed in two ways, either by direct seasonal adjustment of the headline series or by making seasonally adjusted series for a set of underlying series and then aggregate them. We will call the second method the indirect approach of seasonal adjustment.

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<sup>7</sup> Several European countries investigated in the past years what is a better predictor of the expenditures distribution in year  $t-1$ . Was it either the expenditures distribution in year  $t-2$  or was it the expenditures distribution in year  $t-2$  price-updated to the year  $t-1$ ? It came out that the expenditure distribution was the better predictor.

The indirect approach might be preferred, since the seasonal adjustment factors for the separate product groups are easier to identify, estimate and interpret than the seasonal pattern of the headline series which may be composed of various different seasonal patterns. Another relevant aspect is the fact that changing consumption patterns in the course of time may affect the seasonal adjustment factors of the aggregate index, even if the seasonal patterns of the prices for the underlying product groups do not change. Extracting seasonal patterns of underlying series whose profiles are pronounced and sufficiently stable over time may give better results than an adjustment for a changing pattern when applying a direct approach of seasonal adjustment to the headline series. For the purpose of interpretation, inflation analyses and forecasting the indirect approach to seasonal adjustment has the advantage of providing a set of seasonally adjusted component series which perfectly aggregate to the total series. By contrast, a directly seasonally adjusted total series might deviate to a significant extent from the aggregate of its seasonally adjusted component series.

It is to the researchers to decide which of the underlying series are to be seasonally adjusted and for which series the unadjusted series can be used. However, the sum of all the minor seasonalities in the unadjusted series used in the aggregation may add up to some seasonal pattern in the seasonally adjusted aggregate headline series. If a direct approach of the seasonal adjustment of the headline index were used all these minor seasonal effects may also be extracted.

We will not go deeper into this choice between the direct or indirect approach, but for the rest concentrate on how the indirect approach is to be performed.

## 7. Weights in a fixed base seasonally adjusted CPI-series

We will now discuss what the weights will be if we make a seasonal adjustment in a number of the underlying price index series of product groups  $i$ . These were the equations for the unadjusted series:

$$(14) \quad U_{a(m,t)} = \frac{\sum_i P_{i(m,t)} * Q_{ia}}{\sum_i P_{ia} * Q_{ia}} = \sum_i W_{ia} * \frac{P_{i(m,t)}}{P_{ia}}$$

with

$$(15) \quad W_{ia} = \frac{P_{ia} * Q_{ia}}{\sum_i P_{ia} * Q_{ia}}$$

Now if we use seasonally adjusted price series instead of the unadjusted series, these equations become:

$$(16) \quad U_{a(m,t)}^{sa} = \frac{\sum_i P_{i(m,t)}^{sa} * Q_{ia}^{sa}}{\sum_i P_{ia}^{sa} * Q_{ia}^{sa}} = \sum_i W_{ia}^{sa} * \frac{P_{i(m,t)}^{sa}}{P_{ia}^{sa}}$$

with

$$(17) W_{ia}^{sa} = \frac{P_{ia}^{sa} * Q_{ia}^{sa}}{\sum_i P_{ia}^{sa} * Q_{ia}^{sa}}$$

For the compilation of (16) we need both the seasonally adjusted series and the seasonally adjusted weights. The series can be calculated by researchers on the basis of published series, but only the unadjusted weights  $W_{ia}$  are available. The question therefore is under what conditions the weights  $W_{ia}$  and the seasonally adjusted weights  $W_{ia}^{sa}$  are equal.

We can write

$$(18) W_{ia}^{sa} = W_{ia} * \frac{P_{ia}^{sa} * Q_{ia}^{sa}}{\sum_i P_{ia}^{sa} * Q_{ia}^{sa}} * \frac{\sum_i P_{ia} * Q_{ia}}{P_{ia} * Q_{ia}}$$

As explained in section 4 we are dealing only with seasonal pattern of prices and not with a seasonal pattern of consumption. Therefore we can replace  $Q_{ia}^{sa}$  by  $Q_{ia}$  and the equation (18) reduces to

$$(19) W_{ia}^{sa} = W_{ia} * \frac{P_{ia}^{sa}}{P_{ia}} * \frac{\sum_i P_{ia} * Q_{ia}}{\sum_i P_{ia}^{sa} * Q_{ia}}$$

It is clear that if all  $P_{ia}^{sa}$  are equal to  $P_{ia}$  the weights  $W_{ia}^{sa}$  and  $W_{ia}$  are equal.

However, whether or not  $P_{ia}^{sa}$  and  $P_{ia}$  are equal may depend on the procedure that was used for the calculation of the seasonally adjusted price index series and therefore cannot be taken for granted.

We may however conclude that in general the weights  $W_{ia}^{sa}$  and  $W_{ia}$  are equal if

$$\frac{P_{ia}^{sa}}{P_{ia}} * \frac{\sum_i P_{ia} * Q_{ia}}{\sum_i P_{ia}^{sa} * Q_{ia}} = 1.$$

Considering that both  $P_{ia}^{sa}$  and  $P_{ia}$  are annual averages of prices we may assume that, in practice, they will tend to be very close to each other.

## 8. Weights in a chain linked seasonally adjusted CPI-series

In a chain linked Laspeyres index series weights are updated annually and they are determined in a way that differs from the fixed base Laspeyres case. The formula for the weights was developed in section 4 and according to equation (10)

$$(20) \quad W_{i(D,A)} = \frac{P_{i(D,A)} * Q_{i(D,A)}}{\sum_i P_{i(D,A)} * Q_{i(D,A)}}$$

is the December weight for product group  $i$  in the year  $A$ . This weight is used from December  $A$  till December  $A+1$ .

Again we are dealing only with seasonal pattern of prices and not with a seasonal pattern of consumption. Therefore after replacing  $Q_{i(D,A)}$  by  $Q_{iA}$  the equation (20) reduces to

$$(21) \quad W_{i(D,A)} = \frac{P_{i(D,A)} * Q_{iA}}{\sum_i P_{i(D,A)} * Q_{iA}}$$

Likewise the weights for the seasonally adjusted series can be written as:

$$(22) \quad W_{i(D,A)}^{sa} = \frac{P_{i(D,A)}^{sa} * Q_{iA}^{sa}}{\sum_i P_{i(D,A)}^{sa} * Q_{iA}^{sa}}$$

Combining (21) and (22) we can write

$$(23) \quad W_{i(D,A)}^{sa} = W_{i(D,A)} * \frac{P_{i(D,A)} * Q_{iA}^{sa}}{P_{i(D,A)} * Q_{iA}} * \frac{\sum_i P_{i(D,A)} * Q_{iA}}{\sum_i P_{i(D,A)}^{sa} * Q_{iA}^{sa}}$$

The last term of this expression is constant for all product groups  $i$  and also we know that  $Q_{iA}^{sa}$  is equal to  $Q_{iA}$  and that the sum of the weights must be 1. Therefore the expression (23) can be reduced to

$$(24) \quad W_{i(D,A)}^{sa} \propto W_{i(D,A)} * \frac{P_{i(D,A)}^{sa}}{P_{i(D,A)}}$$

and where

$$(25) \quad \sum_i W_{i(D,A)}^{sa} = 1$$

The seasonally adjusted December price  $P_{i(D,A)}^{sa}$  is not equal to the non-adjusted price  $P_{i(D,A)}$  and therefore  $W_{i(D,A)}^{sa}$  is not in general equal to  $W_{i(D,A)}$ .

## 9. Calculating the weights in practice

Having established what the correct weights should be in an indirect approach of seasonal adjustment of the HICP, the question remains how it works out in practice.

First of all the question is whether all necessary information is publicly available. Unfortunately this is not the case.

The weights  $W_{i(D,A)}$  are published by Eurostat. Price indices  $P_{i(D,A)}$  are also published. Eurostat does not perform any seasonal adjustment on any of the series. Researchers who want to perform seasonal adjustments can perform them by themselves on the basis of the published original series to obtain  $P_{i(D,A)}^{sa}$ .

But even then not all necessary data are available for researchers that want to calculate the correct weights and perform the seasonal adjustment. The problem is that the weights for December of year  $A$  are based on the consumption expenditures of the year  $A-1$ , price updated from annual average prices to December prices of this new basket. On the other hand the published price index for December of year  $A$  and the published average annual index for year  $A$  were still calculated on the basis of the basket of year  $A-2$ . Therefore the impact of the price-updating process of the new weights cannot be calculated exactly.

## 10. A practical example; comparison of series

In this section we will compare some results of indirect seasonal adjustment using different weights. Statistics Netherlands does not publish any seasonally adjusted time series for the CPI or HICP of the Netherlands. In order to make a comparison we used time series of seasonally adjusted HICP results for the euro area.

The European Central Bank (ECB) publishes a number of seasonally adjusted series for the euro area HICP. We extracted these series from their database ([www.ecb.int](http://www.ecb.int)). Original HICP series and weights were extracted from the Eurostat database. We used the following series to make test calculations for the aggregation.

The first subdivision is in four major groups of products:

- Food (including alcohol and tobacco),
- Non-energy industrial goods,
- Services and
- Energy. For Energy the ECB does not publish a seasonally adjusted series, and therefore we used the original one<sup>8</sup>.

A summary table 1 with all data and results is at the end of this document.

In the second subdivision we subdivided Food into 5 parts to get a total of 8 series:

- Meat,
- Fish and seafood,
- Fruits,

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<sup>8</sup> According to the ECB's DG Statistics identifiable seasonality is not found in the euro area HICP series for energy.

- Vegetables and
- The other Food subgroups called “Processed food including alcohol and tobacco”.

Results of this subdivision are in summary table 2. The original weights as published and the corrected weights after applying equation (24) are in tables 3 and 4.

To test the possible accuracy of the aggregation process we first calculated aggregated time series from December 2003 till December 2012 from the published original series. Since the published series for the euro area are rounded at two decimal figures (both the indices and the weights) there are minor rounding differences in the calculation.

Secondly we replaced the original index series by the seasonally adjusted series, as published by the ECB and we used the same weights for the unadjusted series.

Finally we corrected the weights according to equation (24) and (25) using the published December figures for the original and seasonally adjusted series, and again performed the aggregation of the seasonally adjusted series.

The results of the calculations can be summarized as follows:

- The level of detail in the publication of the Eurostat data (2 decimal places for indices and for weights) allows the recalculation of the aggregation process with rather high precision. Since 2006 the difference between the published and recalculated figures in table 1 is between -0.0089 and +0.0135. Before 2006 (series 1996=100) the indices were published at one decimal place and the differences are larger. The results in table 2 are almost the same.
- The ECB currently calculates the seasonally adjusted headline inflation figure from aggregation of the four series specified in table 1 using the original weights published by Eurostat. The difference between my recalculation and the ECB publication is between -0.0006 and +0.0004 may be attributable to the rounding at 5 decimal places of the seasonally adjusted series. The differences in table 2 are bigger than in table 1
- After replacing the weights by adjusted weights in the aggregation the indices develop a bit faster. While the differences are zero in 2005 they are 0.03 by the end of 2012, both in table 1 and 2.

It appears that the differences between the aggregated series using unadjusted weights and those using adjusted weights are very small in this practical example. The practical importance in this example is therefore limited. The question remains under what conditions the use of adjusted weights may have a larger impact.

Under what conditions may we expect a higher impact?

- It depends on the size of the seasonal factor in December. If December prices are on the long term trend line the right hand term of equation (24) becomes 1 and the weights are the same.

- It depends on the difference between the price trend of the product with a large seasonal pattern and the overall inflation. Product groups that closely follow the general inflation trend have hardly any impact on the aggregated results. In this respect it is interesting to see that the ECB does not calculate a seasonally adjusted series for energy where the price trend is far above average inflation.

Furthermore the fact that the results were calculated for euro area average figures may have contributed to the small size of the impact. The differences may be larger for some individual countries but cancel out to a larger extent in the euro area aggregate.

Lastly the presented example is an aggregation of four rather high aggregates. It may be that within these aggregates certain seasonal effects at lower level and the impact of the correct weights have cancelled out. The available data did not allow to test this hypothesis.

## A. Glossary

The following notation was used in the text and equations.

### *Variables*

$P$  Price

$Q$  Quantity

$W$  Weight factor

$U$  Laspeyres or Lowe index

### *Subscripts*

$i$  subscript denoting product group

$A$  subscript denoting year  $A$

$(m,A)$  subscript denoting month  $m$  of year  $A$

$(D,A)$  subscript denoting month December of year  $A$

$a$  subscript denoting base period

$b$  subscript denoting reporting period

### *Superscript*

$sa$  superscript denoting seasonally adjusted figure

### *Miscellaneous*

$\propto$  is proportional to

Example:

$U_{2005(m,A)}$  is the index comparing prices in month  $m$  of year  $A$  with the average price level in the year 2005.















**Table 3: HICP Index series, 2005=100, euro area; Weights and corrected weights for SA-series**

COICOP	Original weights; source Eurostat					Price-updated weights for SA-series				
	Food including alcohol and tobacco	Energy (not sa)	Non-energy industrial goods	Services (overall index excluding goods)	All-items HICP	Food including alcohol and tobacco	Energy (not sa)	Non-energy industrial goods	Services (overall index excluding goods)	All-items HICP (before rescaling)
2012	190,71	109,87	284,72	414,70	1.000,00	191,0645	109,87	281,7036	414,1804	996,8185
2011	192,98	103,64	289,07	414,31	1.000,00	193,3511	103,64	286,0829	413,7964	996,8704
2010	191,67	95,59	292,94	419,79	999,99	192,0380	95,59	290,0899	419,2388	996,9567
2009	193,32	95,71	297,24	413,73	1.000,00	193,6870	95,71	294,5774	413,2021	997,1766
2008	195,01	98,15	297,87	408,97	1.000,00	195,3533	98,15	295,4898	408,4511	997,4443
2007	195,57	96,15	300,04	408,23	999,99	195,9317	96,15	297,7391	407,7114	997,5323
2006	192,75	91,99	307,41	407,84	999,99	193,1749	91,99	305,2190	407,2965	997,6804
2005	195,62	85,97	310,32	408,09	1.000,00	196,1675	85,97	308,2796	407,6391	998,0562
2004	195,26	81,32	310,10	413,33	1.000,01	195,9292	81,32	308,2137	412,9488	998,4117

**Table 4: HICP Index series, 2005=100, euro area; Weights and corrected weights for SA-series; 5 food categories**

COICOP	Original weights; source Eurostat					Price-updated weights for SA-series												
	Meat	Fish and seafood	Fruit	Vegetables	Processed food including alcohol and tobacco	Energy (not sa)	Non-energy industrial goods	Services (overall index excluding goods)	All-items HICP	Meat	Fish and seafood	Fruit	Vegetables	Processed food including alcohol and tobacco	Energy (not sa)	Non-energy industrial goods	All-items HICP (before rescaling)	
2012	35,78	10,60	11,76	14,05	118,52	109,87	284,72	414,70	1.000,00	35,7219	10,5706	11,9337	14,1780	118,6171	109,87	281,7036	414,1804	996,7754
2011	35,69	10,41	11,66	15,76	119,46	103,64	289,07	414,31	1.000,00	35,6337	10,3773	11,8398	15,9352	119,5621	103,64	286,0829	413,7964	996,8674
2010	36,08	10,79	11,43	14,66	118,71	95,59	292,94	419,79	999,99	36,0222	10,7500	11,6218	14,8253	118,8064	95,59	290,0899	419,2388	996,9444
2009	36,48	11,07	11,59	15,41	118,77	95,71	297,24	413,73	1.000,00	36,4184	11,0220	11,7882	15,5938	118,8655	95,71	294,5774	413,2021	997,1775
2008	36,76	11,46	12,12	15,70	118,97	98,15	297,87	408,97	1.000,00	36,6938	11,4037	12,3318	15,8652	119,0790	98,15	295,4898	408,4511	997,4645
2007	37,52	11,70	11,71	15,37	119,28	96,15	300,04	408,23	1.000,00	37,4532	11,6416	11,9222	15,5203	119,4206	96,15	297,7391	407,7114	997,5584
2006	36,91	11,55	11,34	14,56	118,40	91,99	307,41	407,84	1.000,00	36,8499	11,4927	11,5677	14,6983	118,5631	91,99	305,2190	407,2965	997,6772
2005	37,77	11,45	11,40	14,76	120,24	85,97	310,32	408,09	1.000,00	37,7225	11,3981	11,6652	14,9354	120,4487	85,97	308,2796	407,6391	998,0586
2004	38,23	11,38	11,77	15,51	118,36	81,32	310,10	413,33	1.000,00	38,1973	11,3304	12,0894	15,7116	118,6072	81,32	308,2137	412,9488	998,4184